

Short Communication: Sensitivity Analysis of an Integrated Climate-Economic Model*

Benjamin M. Bolker[†], Matheus R. Grasselli[‡], and Emma Holmes[§]

Abstract. We conduct a sensitivity analysis of a new type of integrated climate-economic model recently proposed in the literature, where the core economic component is based on the Goodwin–Keen dynamics instead of a neoclassical growth model. Because these models can exhibit much richer behavior, including multiple equilibria, runaway trajectories, and unbounded oscillations, it is crucial to determine their sensitivity to changes in underlying parameters. We focus on four economic parameters (markup rate, speed of price adjustments, coefficient of money illusion, growth rate of productivity) and two climate parameters (size of upper ocean reservoir, equilibrium climate sensitivity) and show how their relative effects on the outcomes of the model can be quantified by methods that can be applied to an arbitrary number of parameters.

Key words. climate change, integrated assessment models, stock-flow consistent models, ecological economics

AMS subject classifications. 91B55, 91B76, 34C60

DOI. 10.1137/21M1404120

1. Introduction. Climate change is recognized as one of the largest risks facing the global financial system. Losses due to extreme weather events alone have risen tenfold in the past 40 years, with a 10-year average now over \$200 billion per year. Even the transition to a low-carbon economy poses challenging risks, if only because of the size of the dislocation from carbon-intensive portfolios, with total pledged divestment approaching \$15 trillion worldwide. Conversely, the needed investment in green technology, mitigation, and infrastructure is at least an order of magnitude larger than current investment flows, thus presenting a growth opportunity for innovative green finance initiatives.¹ To adequately address these risks, financial mathematicians need to use and develop models that integrate economic and climate dynamics in a coherent framework. Mainstream models such as the dynamic integrated climate-economic (DICE) model [12] and its variants make a poor foundation; not only do they generally omit

*Received by the editors March 10, 2021; accepted for publication (in revised form) May 7, 2021; published electronically June 21, 2021.

<https://doi.org/10.1137/21M1404120>

Funding: This work was funded by a Discovery Grant from the Natural Sciences and Engineering Research Council (NSERC) of Canada.

[†]Departments of Mathematics & Statistics and Biology, McMaster University, Hamilton, ON L8S 4K1, Canada (bolker@math.mcmaster.ca).

[‡]Department of Mathematics & Statistics, McMaster University, Hamilton, ON L8S 4K1, Canada (grasselli@math.mcmaster.ca).

[§]Department of Statistical Sciences, University of Toronto, Toronto, ON M5G 1Z5, Canada (emmarobin.holmes@mail.utoronto.ca).

¹For extreme-weather losses, see the Sigma 2 Report by Swiss Re (www.swissre.com). For divestment commitments, see Fossil Free (<https://gofossilfree.org/divestment/commitments/>). For green investment needs and current flows, see the Global Landscape of Climate Finance 2019 Report (www.climatepolicyinitiative.org).

SC44

a banking or financial sector, but they are also methodologically incompatible with the most salient features of climate science, such as nonlinear feedback mechanisms and tipping points [10]. A better alternative consists of stock-flow consistent models, in which both the economy and the climate are modeled as a system of nonlinear differential equations describing slowly adjusting, out-of-equilibrium quantities [2]. Such models do *not* assume an equilibrium path with a growing economy a priori. The dynamics of key economic variables can exhibit runaways, unbounded oscillations, and convergence to undesirable equilibria, similar to the behavior of climate variables in some regimes.

On the other hand, the output dynamics of these models can be very sensitive with respect to parameter values. In particular, due to the possibility of multiple equilibria, small changes in parameters or initial values can lead to completely different long-term values for key economic variables. A preliminary sensitivity analysis of the model in [2] was conducted in [3] in order to investigate the effect of uncertainty in productivity growth, equilibrium temperature sensitivity, and a carbon absorption parameter. The purpose of this paper is to extend this analysis in several significant ways.

We start by considering the sensitivity of the model with respect to three additional economic parameters: the speed of price adjustment to fluctuations, the markup rate over labor costs used by firms, and the degree of money illusion in wage bargaining. We establish that the sensitivity with respect to the markup rate is particularly important, because the model converges to two very different equilibria as this parameter varies within the range of empirically observed values.

Based on this observation, we conduct a more detailed analysis to quantify the influence of different parameters on the likelihood of the model converging to an interior equilibrium or one exhibiting explosive behavior. Specifically, we perform a logistic regression against the several model parameters, with the categorical response variable describing whether the long-term employment rate remains above a given threshold. In this way, we are able to confirm that the markup rate has the largest influence on whether the model converges to an equilibrium with low or high employment in the long run, and that this effect persists when a full feedback from climate change, in terms of both damages and policy responses, is taken into account.

Finally, whereas [3] reports the distribution of some key output variables, such as temperature anomaly and private debt ratio, when parameter values are drawn from their own distributions, it did not describe which parameters contributed the most to the variation in the output. To address this, we use the technique adopted in [1] and compute the partial rank correlation of each parameter under consideration with the employment in year 2100, conditional on it being above a threshold (that is to say, conditional on a “good” equilibrium). Apart from establishing which parameters are positively or negatively correlated with the employment rate, our results indicate that the magnitude of the effect of uncertainty in economic parameters on the model outputs is comparable with that of uncertainty in climate parameters.

2. The model. We describe the core economic model without climate change first, followed by the full model with climate damages and policy responses. In what follows, constant parameters are denoted with a bar, whereas a dot denotes a time derivative.

2.1. The economy. We adopt the formulation presented in [7], based on the original model proposed in [9], with the necessary modifications to make the model compatible with [2] and [3]. The model makes three key sets of assumptions. The first concerns the relation of output with capital and labor in the economy. We assume that real (i.e., inflation adjusted) output is given by $Y = \frac{K}{\bar{\nu}}$, where $\bar{\nu}$ is a constant capital-to-output ratio and K is the real capital stock, which evolves according to $\dot{K} = I - \bar{\delta}K$. Here I denotes real investment by firms and $\bar{\delta}$ is a constant depreciation rate. From real output Y , we can obtain the number of employed workers $L = \frac{Y}{a}$, where a denotes the productivity per worker. Denoting the total workforce by N , it follows that the employment rate is given by $\lambda = \frac{L}{N} = \frac{Y}{aN}$.

The second set of assumptions has to do with the behavior of firms. Denote nominal profits by $\Pi = pY - wL - \bar{r}D$, where pY is the total sales revenue of real output Y at a price level p , w is the average nominal wage rate per worker, and \bar{r} denotes an average constant rate of interest paid on net debt D . The model assumes that real investment by firms is given by $I = \kappa(\pi)Y$ for a function $\kappa(\cdot)$ of the profit share of nominal output

$$(2.1) \quad \pi = \frac{\Pi}{pY} = 1 - \omega - \bar{r}d,$$

where we have introduced the wage share $\omega = \frac{wL}{pY}$ and the debt-to-output ratio $d = \frac{D}{pY}$. In the absence of any other source of financing, firms have to fund this investment by either using profits or borrowing from banks, from which it follows that the change in net debt of firms is given by $\dot{D} = pI - \Pi + \Delta(\pi)pY$, where the last term denotes dividends paid to shareholders.²

The final set of assumptions corresponds to the determination of wages and prices. We assume that the wage rate changes according to $\frac{\dot{w}}{w} = \Phi(\lambda) + \bar{\gamma}i(\omega)$, where $\Phi(\cdot)$, known as the Phillips curve, represents the bargaining power of workers as a function of the employment rate; $\bar{\gamma} \geq 0$ is a coefficient measuring the degree of money illusion (with no illusion corresponding to $\bar{\gamma} = 1$); and $i(\omega)$ corresponds to the inflation rate, which is assumed to be of the form³

$$(2.2) \quad i(\omega) = \frac{\dot{p}}{p} = \bar{\eta}(\bar{\xi}\omega - 1)$$

for an adjustment parameter $\bar{\eta} > 0$ and a markup factor $\bar{\xi} \geq 1$.

Finally, we make two additional assumptions that can be relaxed without altering the model in any significant way, namely, that labor productivity a grows exponentially at a constant rate $\bar{\alpha}$ and that the workforce N follows the sigmoid function expressed below in (2.3d). With the assumptions and definitions in place so far, the economy can be described by the following four-dimensional system⁴ of coupled nonlinear differential equations for the

²The original Keen model in [9] does not use dividends, but [2] found it necessary to add this term to the model in order to improve the empirical estimates.

³The inflation function used in [2] includes the cost of capital and carbon taxes, in addition to labor, as a cost of production for firms. In [3], this is dropped in favor of the simpler inflation dynamics adopted here. Accordingly, the parameters $(\bar{\eta}, \bar{\xi})$ changed from (0.5, 1.3) in [2] to (0.192, 1.875) in [3]. Both [2] and [3] assume that $\bar{\gamma} = 0$.

⁴In the original model in [9], the growth rate of N is assumed to be a constant $\bar{\beta}$, so that the relevant dynamics reduces to a three-dimensional system with state variables (λ, ω, d) .

state variables (λ, ω, d, N) :

$$(2.3a) \quad \frac{\dot{\lambda}}{\lambda} = \frac{\kappa(\pi)}{\bar{\nu}} - \bar{\delta} - \bar{\alpha} - \bar{\delta}_N \left(1 - \frac{N}{\bar{N}_{\max}}\right),$$

$$(2.3b) \quad \frac{\dot{\omega}}{\omega} = \Phi(\lambda) - \bar{\alpha} - (1 - \bar{\gamma})i(\omega),$$

$$(2.3c) \quad \frac{\dot{d}}{d} = \frac{\kappa(\pi) - \pi + \Delta(\pi)}{d} - \left[i(\omega) + \frac{\kappa(\pi)}{\bar{\nu}} - \bar{\delta} \right],$$

$$(2.3d) \quad \frac{\dot{N}}{N} = \bar{\delta}_N \left(1 - \frac{N}{\bar{N}_{\max}}\right),$$

where π is defined in (2.1) and $\kappa(\cdot)$, $\Phi(\cdot)$, and $\Delta(\cdot)$ are functions that need to be calibrated. For concreteness, we use a linear function for the Phillips curve⁵ and truncated linear functions for the investment and dividends, with parameters specified in Table 1.

A full analysis of the equilibria for (2.3) is presented in [7] and summarized here. The interior equilibrium, corresponding to a desirable economic situation of nonzero wages and employment, is given by

$$(2.4) \quad (\lambda^*, \omega^*, d^*, N^*) = \left(\Phi^{-1}(\bar{\alpha} + (1 - \bar{\gamma})i(\omega^*)), 1 - \pi^* - rd^*, \frac{\kappa(\pi^*) - \pi^* + \Delta(\pi^*)}{\bar{\alpha} + i(\omega^*)}, \bar{N}_{\max} \right),$$

where $\pi^* = \kappa^{-1}[\bar{\nu}(\bar{\alpha} + \bar{\delta})]$. As we can see, substituting the expression for d^* into the expression for ω^* leads to a quadratic equation, from which one can deduce conditions for the existence of at least one strictly positive solution (and sometimes two). There are three other possible equilibria for (2.3), all of which correspond to undesirable economic outcomes. The first corresponds to the state variables (λ, ω, d) converging to $(0, 0, \infty)$, namely, vanishing wage share and employment rate and an explosive debt ratio, as first identified in [5]. The second and third represent equilibrium outcomes where the employment rate is zero but the wage share is positive and given by $\omega^{**} = \frac{1}{\xi} + \frac{\Phi(0) - \bar{\alpha}}{\xi\bar{\eta}(1 - \bar{\gamma})}$, in which case the equilibrium debt share can be either finite or infinite.

2.2. The climate. The climate part of the model follows [2], which uses a continuous-time version of the DICE model [12]. It begins by specifying the amount of carbon emissions associated with a level Y^0 of industrial production as $E_{ind} = \sigma(1 - n)Y^0$, where σ is the carbon intensity of the economy and n is an emissions reduction rate. Regardless of the decisions of firms, we assume that technological progress gradually leads the carbon intensity σ to decrease in time with a rate $g_\sigma < 0$, which in turn approaches zero at a constant rate $\bar{\delta}_{g_\sigma} < 0$.

To accelerate the transition to an emission-free economy, firms can choose a reduction rate n , for which they have to pay abatement costs per unit of production assumed to be of the form $A = \frac{\sigma p_{BS} n^\theta}{\theta}$, where the parameter $\bar{\theta} > 0$ controls the convexity of the cost and p_{BS} is the (inflation adjusted) price of a backstop technology, which we assume to decrease exponentially

⁵For comparison, the parameter values for the Phillips curve were chosen to match those in [3], which were estimated under the assumption that $\bar{\gamma} = 0$. These values would be different if $\bar{\gamma}$ were also estimated from data.

Table 1
Model parameters.

Symbol	Value	Parameter description
$\bar{\alpha}$	0.02	Productivity growth rate
$\bar{\delta}$	0.04	Depreciation rate of capital
$\bar{\nu}$	2.7	Capital-to-output ratio
$\bar{\delta}_N$	0.031	Workforce growth parameter
\bar{N}_{max}	7.065	Workforce equilibrium value
$\bar{\Phi}_0$	-0.292	Phillips curve y-intercept
$\bar{\Phi}_1$	0.469	Phillips curve slope
$\bar{\kappa}_0$	0.0318	Investment function y-intercept
$\bar{\kappa}_1$	0.575	Investment function slope
$\bar{\kappa}_{min}$	0	Investment function minimum
$\bar{\kappa}_{max}$	0.3	Investment function maximum
$\bar{\Delta}_0$	-0.078	Dividend function y-intercept
$\bar{\Delta}_1$	0.553	Dividend function slope
$\bar{\Delta}_{min}$	0	Dividend function minimum
$\bar{\Delta}_{max}$	0.3	Dividend function maximum
\bar{r}	0.02	Long-term interest rate
$\bar{\eta}$	0.192	Inflation relaxation parameter
$\bar{\xi}$	1.875	Price markup
$\bar{\gamma}$	0.9	Effect of inflation on wages
\bar{C}^{AT}	588	Preindustrial concentration of CO ₂ in the atmosphere
\bar{C}^{UP}	360	Preindustrial concentration of CO ₂ in the upper ocean
\bar{C}^{LO}	1720	Preindustrial concentration of CO ₂ in the lower ocean
$\bar{\phi}_{12}$	0.024	Transfer coefficient for carbon from AT to UP
$\bar{\phi}_{23}$	0.001	Transfer coefficient for carbon from UP to LO
$\bar{\delta}_{g\sigma}$	-0.001	Variation rate of the growth of emission intensity
$\bar{\delta}_{E_{land}}$	-0.022	Growth rate of land use change CO ₂ -equivalent emissions
$\bar{\delta}_{p_{BS}}$	-0.005	Growth rate of the price of backstop technology
\bar{F}_{dbl}	3.681	Change in radiative forcing from a doubling of preindustrial CO ₂
\bar{F}_{exo}^{start}	0.5	Starting value of exogenous radiative forcing
\bar{F}_{exo}^{end}	1	Ending value of exogenous radiative forcing
\bar{T}_{preind}	13.74	Preindustrial temperature, in degrees Celsius
\bar{c}	10.20	Heat capacity of atmosphere and upper ocean layer
\bar{c}^{LO}	3.52	Heat capacity of the lower ocean layer
\bar{h}	0.0176	Heat exchange coefficient between temperature layers
\bar{S}	3.1	Equilibrium climate sensitivity, in degrees Celsius
$\bar{\zeta}_1$	0	Damage function parameter
$\bar{\zeta}_2$	0.00236	Damage function parameter
$\bar{\zeta}_3$	0	Damage function parameter
$\bar{\zeta}_4$	0	Damage function parameter
$\bar{\theta}$	2.6	Abatement cost function parameter
\bar{s}_A	0.5	Fraction of abatement costs subsidized by government
$\bar{\delta}_{C,1}$	25.69	Linear growth rate of the carbon price up to 2020
$\bar{\delta}_{C,2}$	2.53	Linear growth rate of the carbon price from 2020 onwards

at a constant rate $\bar{\delta}_{pBS}$. The incentive to pay this abatement cost comes from the fact that firms are assumed to face a carbon tax of the form $T^C = p_C E_{ind}$, where p_C is the (inflation adjusted) carbon price.⁶ On the other hand, firms also receive a subsidy $S^C = \bar{s}_A AY^0$ equal to a fraction $0 \leq \bar{s}_A < 1$ of their abatement costs. Accordingly, the emission reduction rate that minimizes the sum of carbon tax and abatement cost is given by

$$(2.5) \quad n = \min \left\{ \left(\frac{p_C}{(1 - \bar{s}_A)p_{BS}} \right)^{\frac{1}{\theta-1}}, 1 \right\}.$$

In addition to industrial emissions, the model assumes that there are land-use emissions E_{land} that decrease at a constant rate $\bar{\delta}_{E_{land}} < 0$, so that total emissions are given by $E_T = E_{land} + E_{ind}$. These emissions change the average concentrations of carbon dioxide according to the following three-layer model for the atmosphere, the upper ocean and biosphere, and the lower ocean:

$$(2.6) \quad \begin{pmatrix} \dot{C}O_2^{AT} \\ \dot{C}O_2^{UP} \\ \dot{C}O_2^{LO} \end{pmatrix} = \begin{pmatrix} E_T \\ 0 \\ 0 \end{pmatrix} + \begin{pmatrix} -\bar{\phi}_{12} & \bar{\phi}_{12}\bar{C}_{UP}^{AT} & 0 \\ \bar{\phi}_{12} & -\bar{\phi}_{12}\bar{C}_{UP}^{AT} - \bar{\phi}_{23} & \bar{\phi}_{23}\bar{C}_{LO}^{UP} \\ 0 & \bar{\phi}_{23} & -\bar{\phi}_{23}\bar{C}_{LO}^{UP} \end{pmatrix} \begin{pmatrix} C O_2^{AT} \\ C O_2^{UP} \\ C O_2^{LO} \end{pmatrix},$$

where $\bar{C}_j^i = \frac{\bar{C}^i}{\bar{C}^j}$ for $i, j = \{AT, UP, LO\}$ are ratios of preindustrial concentrations \bar{C}^{AT} , \bar{C}^{UP} , and \bar{C}^{LO} and $\bar{\phi}_{12}$, $\bar{\phi}_{23}$ are transfer parameters. In turn, an increase in atmospheric CO₂ concentration leads to an increase in the Earth’s radiative forcing, namely, the influx of solar energy that is not radiated back into space. Specifically, the model assumes that the increase in radiative forcing F_{ind} caused by industrial emissions is a linear function of the logarithm of the ratio of CO₂^{AT} to its preindustrial level. The planet’s total increase in radiative forcing from preindustrial levels is then given by $F = F_{ind} + F_{exo}$, where F_{exo} is an exogenous increasing function that approaches a constant. Finally, the radiative forcing affects the interplay between the temperature anomaly (compared to preindustrial levels) T for the atmosphere and upper ocean and the corresponding temperature anomaly T^{LO} for the lower ocean according to the dynamics

$$(2.7a) \quad \bar{c}\dot{T} = F - \frac{\bar{F}_{dbl}}{\bar{S}}T - \bar{h}(T - T^{LO}),$$

$$(2.7b) \quad \bar{c}^{LO}\dot{T}^{LO} = \bar{h}(T - T^{LO}),$$

where \bar{F}_{dbl} is a parameter that represents the increase in radiative forcing caused by doubling of preindustrial atmospheric CO₂ concentration, \bar{c} and \bar{c}^{LO} are the heat capacities of each layer, and \bar{h} is a parameter representing the heat exchange between layers. It follows from (2.7a)–(2.7b) that the equilibrium climate sensitivity (ECS)—defined as the equilibrium temperature anomaly resulting from a doubling of preindustrial atmospheric CO₂ concentration—is given by the parameter \bar{S} .

⁶For direct comparison with [2] and [3], we assume that the carbon price follows the upper bound of the Stern–Stiglitz corridor, namely, reaching \$80/tC by 2020 and \$100/tC by 2030 in 2005 USD units.

To summarize, in the absence of any further feedback to the economy, the climate part of the model can be viewed as a component that takes production Y^0 as an input and returns the average increase in global temperature T as an output through the following chain of causal relationships explained above: $Y^0 \rightarrow E_{ind} \rightarrow E_T \rightarrow \text{CO}_2^{AT} \rightarrow F_{ind} \rightarrow F \rightarrow T$.

2.3. Coupling. The economy and the climate are coupled in a unified model as follows. First, as mentioned in the previous section, for a given level of production $Y^0 = \frac{K}{\bar{v}}$, firms incur abatement costs AY^0 in order to achieve an emissions reduction rate n . These costs are subtracted directly from production, so that only $(1 - A)Y^0$ is available as output for sale. Next, because of climate change, a fraction \mathbf{D} of this output is assumed to be irreparably damaged. This fraction is assumed to be a function of temperature anomaly T :

$$(2.8) \quad \mathbf{D} = 1 - \frac{1}{1 + \bar{\zeta}_1 T + \bar{\zeta}_2 T^2 + \bar{\zeta}_3 T \bar{\zeta}_4},$$

where $\bar{\zeta}_1, \bar{\zeta}_2, \bar{\zeta}_3, \bar{\zeta}_4$ are given parameters.⁷ Consequently, the output actually sold by firms is given by $Y = (1 - \mathbf{D})(1 - A)Y^0$. Accordingly, profits for firms need to be modified as

$$(2.9) \quad \Pi = (1 - \mathbf{D})(1 - A)pY^0 - wL - \bar{r}D + p(S^C - T^C),$$

where S^C and T^C are the (real) government subsidy and carbon tax mentioned in the previous section, with the correspondingly redefined profit share $\pi = \frac{\Pi}{pY}$.

In order to avoid taking derivatives of the damage function and abatement costs, it is computationally more convenient to use the extensive variables (K, D, w, p, a, N) as state variables for the economic model, instead of the intensive-form system in terms of the ratios (λ, ω, d) , although both these formulations can be shown to be equivalent. Adding the state variables $(\sigma, g_\sigma, E_{land}, \text{CO}_2^{AT}, \text{CO}_2^{UP}, \text{CO}_2^{LO}, T, T^{LO}, p_{BS}, p_C)$ from the climate component described in the previous section leads to a 16-dimensional combined climate-economic model.

A limited analysis of the equilibria for the full model is provided in [2]. Assuming no inflation, they show that once the temperature level has reached equilibrium, the “good” equilibrium—where the economy grows at a constant rate, employment and wages are positive, and the debt ratio is finite—exists. This equilibrium is similar to (2.4). Further scenario analysis is done numerically in [2], and a sensitivity analysis with respect to the parameters $\bar{\alpha}$ (growth rate of output), \bar{S} (equilibrium climate sensitivity), and \bar{C}^{UP} (size of the intermediate climate reservoir) is provided in [3]. In the next section, we present a more complete sensitivity analysis taking into account some key economic parameters.

3. Sensitivity analysis. We investigate the sensitivity of the economic model without climate change first, followed by an analysis of the full model. Numerical results were obtained by solving the models in R using the package `deSolve` [15] with the `lsoda` integration method.⁸

⁷The desired convexity of such damage curves is contested in the literature: see [4] for critiques of the function used in [12]. Nevertheless, for comparison with [2] and [3] we use the Nordhaus damage function as in [12]. This assumption can be relaxed to allow for varying, and arguably more realistic, levels of damages.

⁸The code used in this paper is available at <https://github.com/emmaholmes/econ-climate-sensitivity> and is an extension of the code provided by the authors of [2] and [3]. Any remaining errors are ours.

3.1. Sensitivity of the economic model. As pointed out in [7], depending on the choice of parameters, the pricing dynamics (2.2) can significantly alter the outcome of the underlying Keen model, because of both the possibility of deflation and the introduction of new undesirable equilibrium points.

To explore the parameter space, the model was run up to year 2300 as in [2], with 20 output time-steps per year, for inflation parameters in the ranges shown in Figure 1, which were chosen to include the available empirical estimates. Points were selected from the parameter space using Sobol' sequences [14] implemented by the R package `qrmg` [8].

The outcome of each individual model run was categorized into three possibilities, based on the ending values of the simulation: “good” if $0.4 \leq \lambda \leq 1$, $0.4 \leq \omega \leq 1$, and $d \leq 2.7 = \bar{v}$, corresponding to an economy with employment and wages bounded away from zero and debt less than the total capital stock, that is to say, the “convergence set” specified in footnote 30 of [2]; “outside bounds” if one of λ or ω ended above 1; or “bad” otherwise. The bad outcomes include all of the equilibria corresponding to the vanishing employment rates mentioned above, as well as the interior equilibria with low but positive wage shares, which are associated with low inflation (and sometimes even deflation) through (2.2), and consequently low employment rates through (2.4).

The top graphs in Figure 1 show the results for the economic model alone, starting with two sets of initial conditions: (a) favorable initial conditions, meaning that initially the employment rate and wage share in the economy are high and the debt share is low, and (b) unfavorable initial conditions, with lower initial wages and employment and higher initial debt, matching the initial conditions used in [2]. As we can see in the figure, the key parameter affecting the model outcome is the markup rate $\bar{\xi}$. There are slightly more bad outcomes for low values of the parameter $\bar{\gamma}$ (higher degree of money illusion) or a high relaxation parameter $\bar{\eta}$ (faster price adjustments), but the effects are not pronounced. We see that some choices of pricing parameters create high oscillation in the model or otherwise keep the model from converging to an economically meaningful equilibrium by 2300, as represented by the purple area in the graphs.

An alternative way to explore this result is to look at how changing the markup rate $\bar{\xi}$ changes the basin of attraction to the “good” equilibrium. The graphs (a) to (c) in Figure 2 show the results of running the economic model for a range of initial conditions corresponding to the “initial set” specified in footnote 29 of [2]. The other pricing parameters are fixed at $\bar{\gamma} = 0.9$ and $\bar{\eta} = 0.4$, and we use the same categorized outcomes as in Figure 1. We can see that increasing the markup rate from $\bar{\xi} = 1.3$ (as used in [2] for an inflation specification that included the cost of capital) to $\bar{\xi} = 1.875$ (as used in [3] for the same inflation specification adopted here) increases the basin of attraction to the good equilibrium (2.4), whereas reducing the markup rate to 1.18 has the opposite effect.

For a more detailed analysis, we follow [3] and perform simulations of the model with parameters drawn from probability distributions fitted to the empirical estimates. Namely, the productivity growth rate $\bar{\alpha}$ is drawn from a normal distribution with mean 2.06 and standard deviation 1.12 as in [3], whereas for the inflation parameters we use the estimates in [6, Online Appendix] and assume that (1) $\bar{\eta}$ is drawn from a normal distribution with mean 0.4 and standard deviation of 0.12; (2) $\bar{\xi}$ is drawn from a generalized Gamma distribution with shape parameter $s = 3.0894$, scale parameter $m = 0.7154$, and family parameter $f = 0.9959$,

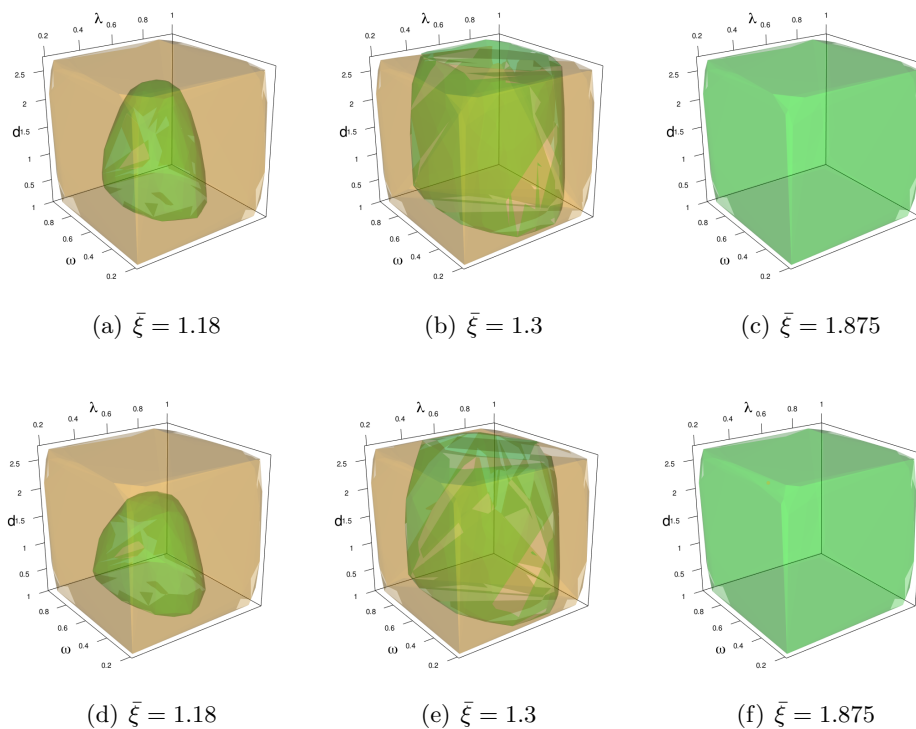


Figure 2. Outcomes of the economic model (top) and the full model (bottom) for a range of initial conditions and three different values of the markup rate $\bar{\xi}$. Increasing the markup rate increases the numerically computed basin of attraction to the “good” equilibrium (green region), confirming the results shown in Figure 1.

specified in [2]. We then sorted the outcome of the model into two categories—good and bad—depending on whether or not the employment rate in 2100 is above 40%, and performed a logistic regression of this categorical outcome with respect to the four parameters, with inputs standardized as described in [13] using the R function `scale`. Next, to quantify how the uncertainty in an input variable affects the output, we followed [1] and computed the partial rank correlation coefficient (PRCC) associated with each parameter, using the employment rate for good outcomes as the output variable. The results are shown in the top panels of Figure 3 (green dots) and indicate that the markup rate has a strong effect on the outcome, confirming the conclusions obtained from the top part of Figure 1. Moreover, conditioned on converging to a good outcome, the markup rate $\bar{\xi}$ and the money illusion parameter $\bar{\gamma}$ have similar effects on the outcome, but in opposite directions. The relaxation parameter $\bar{\eta}$ and the productivity growth rate $\bar{\alpha}$ have slightly smaller effects.

3.2. Sensitivity of the full model. We now repeat the analysis using the full 16-dimensional integrated climate-economic model. As a preliminary result, graphs (c) and (d) in Figure 1 and graphs (d) to (f) in Figure 2 show the outcomes of the full model using the same sets of initial conditions for (λ, ω, d) , parameter values $(\bar{\eta}, \bar{\xi}, \bar{\gamma})$, and classification criterion

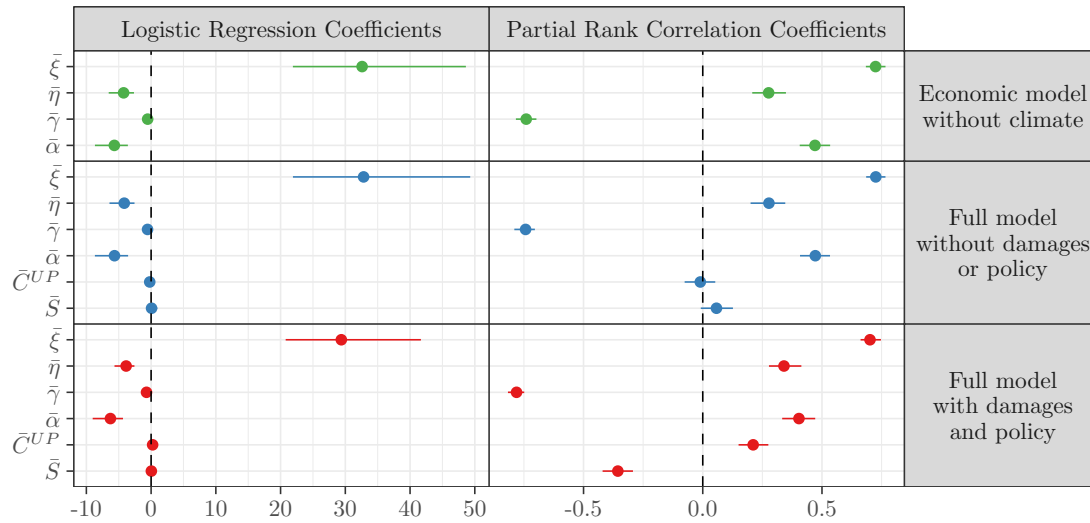


Figure 3. Logistic regression coefficients and partial rank correlation coefficients for the markup rate $\bar{\xi}$, the relaxation parameter $\bar{\eta}$, the money illusion parameter $\bar{\gamma}$, the size of the intermediate climate reservoir \bar{C}^{UP} , and the equilibrium climate sensitivity \bar{S} .

as before, whereas for the remaining initial conditions we use the values specified in [2]. We observe broadly similar results as for the economic model, indicating that the inflation parameters, in particular the markup factor $\bar{\xi}$, still play an important role in determining the long-term behavior of the outcomes in the full model.

For a more detailed sensitivity analysis, we again draw model parameters from probability distributions matching available empirical estimates. In addition to the inflation and growth parameters, we follow [3] and consider uncertainty in two climate-related parameters: the equilibrium climate sensitivity \bar{S} and the size \bar{C}^{UP} of the intermediate climate reservoir. Specifically, \bar{S} is drawn from a log-normal distribution with log-mean 1.107 and log-standard deviation 0.264, and \bar{C}^{UP} is drawn from a log-normal distribution with log-mean 5.886 and log-standard deviation 0.251, which correspond to the distributions used in [3] to match the estimates provided in [11].

As before, the parameters are sampled from their distributions 1000 times and the model is run with the same initial conditions taken from [2]. The same procedure is performed to calculate logistic regression coefficients, where the outcome variable is whether or not the employment rate in 2100 is above 40%, and PRCCs, where the outcome variable is the employment rate in 2100, conditional on being above 40%. The analysis is performed twice, first with the full model but without climate damages and policy, and then again with climate damages, carbon tax, and a government subsidy.

The results are shown in the remaining panels of Figure 3. When neither damages nor government policy are taken into account (blue dots, middle panels), the logistic regression coefficients again indicate that the markup rate $\bar{\xi}$ has the largest effect, with the labor productivity growth rate $\bar{\alpha}$ and the relaxation parameter $\bar{\eta}$ also having effects significantly different

from zero. The PRCC values show that, in this case, uncertainty in the pricing parameters has a greater effect on the outcome than uncertainty in the variables examined in [3], which is unsurprising, given that in this example there is no feedback from the climate into the economic model.

For the full model with climate damages, a carbon tax policy, and government subsidy (red dots, bottom panels), we can see that the climate parameters all have a larger effect, both in determining whether or not the model converges to a good output and in the variability of employment rate in the good outcomes. Interestingly, however, the effects of the inflation parameters are comparable in size to those of the climate ones, meaning that uncertainty in both sets of parameters needs to be taken into account in evaluating the robustness of the model.

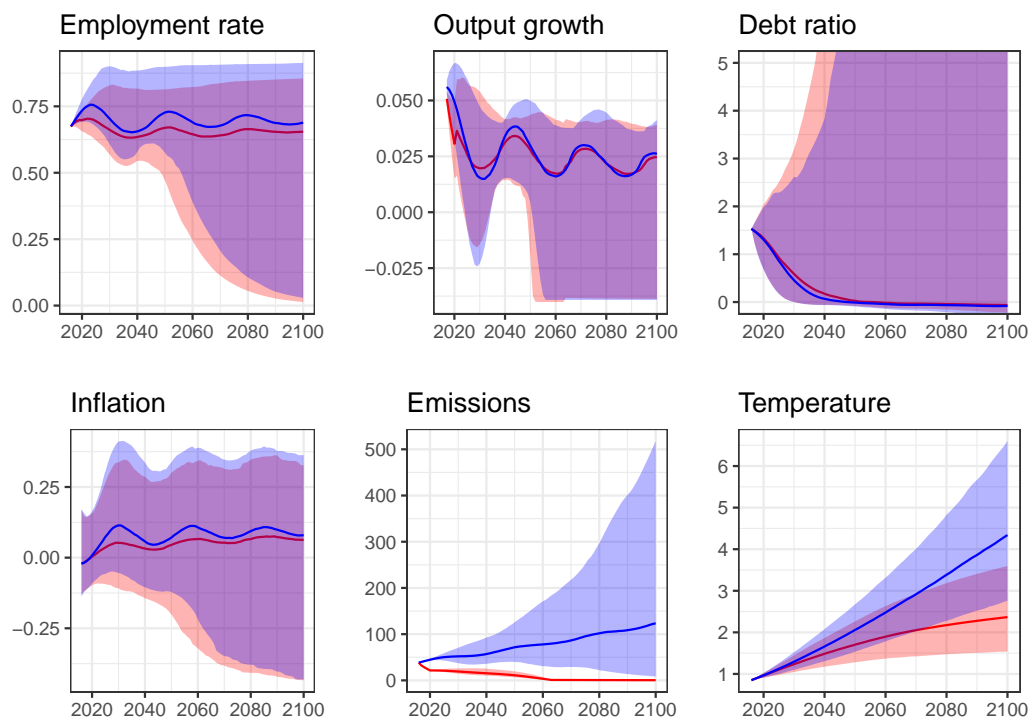


Figure 4. Results of a Monte Carlo simulation of the full model with (in red) and without (in blue) climate damages and government policy. The medians and the 95% confidence intervals are shown for all runs, including both “good” and “bad” outcomes.

As a final illustration, Figure 4 depicts the trajectories of select state variables of the model. Differently from the results presented in Figure 3, this figure shows the values for all runs, instead of only those converging to outcomes with employment higher than 40%, as this allows us to obtain estimates for the unconditional distribution of key state variables. For example, we find that in the full model with climate feedback and government policy, the temperature anomaly by 2100 remains below the 2°C Paris accord target in 23.8% of the runs with a median value 2.37°C, the debt-to-output ratio remains below the acceptable value 2.7

in 84.7% of the runs, and both variables remain below these thresholds in 20% of the runs, broadly similar to the results reported in [3].

4. Conclusion. In this paper, sensitivity analysis was conducted on a Keen-based model of the climate and the economy. The model is similar to that of [2], but with the pricing dynamics of [7] and slight simplifications to the damage curve and path of the carbon price. The parameter space was explored numerically using Sobol' sampling and Monte Carlo methods. We show that convergence of the model to an interior equilibrium is sensitive to small changes in some key parameters: the inflation markup rate $\bar{\xi}$, the labor productivity growth rate $\bar{\alpha}$, and the speed of price adjustment $\bar{\eta}$. Furthermore, conditional on convergence, we show that uncertainty on economic and climate parameters has quantitatively comparable effects on the outcome of the model.

These results indicate that sensitivity analyses of integrated climate-economic models need to take into account uncertainties in *all* relevant parameters, as they can have significant effects on the conclusions drawn from the model. Accordingly, a future avenue for research would be to estimate these parameter values from existing economic data, as was done in [6], but for this specific version of the Keen model with climate and some of its extensions. Such a study would allow for stronger conclusions about the effectiveness of differing government climate policies, such as carbon taxes, green technology subsidies, as well as traditional monetary or fiscal policies, in stabilizing the inherently complex interactions between the climate and the economy.

Acknowledgments. We are grateful for stimulating discussions with Florent McIsaac, Etienne Espagne, Devrim Yilmaz, Antoine Godin, Gaël Giraud, and the participants of the Georgetown Environmental Justice Program Seminar, where this work was presented, and also for the helpful remarks by the anonymous referee and review editor for this paper.

REFERENCES

- [1] S. M. BLOWER, D. HARTEL, H. DOWLATABADI, R. M. ANDERSON, AND R. M. MAY, *Drugs, sex and HIV: A mathematical model for New York City*, Philos. Trans. Roy. Soc. London Ser. B, 331 (1991), pp. 171–187.
- [2] E. BOVARI, G. GIRAUD, AND F. MC ISAAC, *Coping with collapse: A stock-flow consistent monetary macrodynamics of global warming*, Ecol. Econ., 147 (2018), pp. 383–398.
- [3] E. BOVARI, O. LECUYER, AND F. MC ISAAC, *Debt and damages: What are the chances of staying under the 2°C warming threshold?*, Internat. Econ., 155 (2018), pp. 92–108.
- [4] S. DIETZ AND N. STERN, *Endogenous growth, convexity of damage and climate risk: How Nordhaus' framework supports deep cuts in carbon emissions*, Econ. J., 125 (2015), pp. 574–620.
- [5] M. R. GRASELLI AND B. COSTA LIMA, *An analysis of the Keen model for credit expansion, asset price bubbles and financial fragility*, Math. Financ. Econ., 6 (2012), pp. 191–210, <https://doi.org/10.1007/s11579-012-0071-8>.
- [6] M. R. GRASELLI AND A. MAHESHWARI, *Testing a Goodwin model with general capital accumulation rate*, Metroecon., 69 (2018), pp. 619–643.
- [7] M. R. GRASELLI AND A. NGUYEN HUU, *Inflation and speculation in a dynamic macroeconomic model*, J. Risk Financ. Manag., 8 (2015), pp. 285–310, <https://doi.org/10.3390/jrfm8030285>.
- [8] M. HOFERT AND C. LEMIEUX, *qrng: (Randomized) Quasi-Random Number Generators*, R package version 0.0-7, 2019, <https://CRAN.R-project.org/package=qrng>.

- [9] S. KEEN, *Finance and economic breakdown: Modeling Minsky's "Financial instability hypothesis,"* J. Post Keynesian Econ., 17 (1995), pp. 607–635, <http://www.jstor.org/stable/4538470>.
- [10] S. KEEN, *The appallingly bad neoclassical economics of climate change,* Globalizations, (2020), pp. 1–29, <https://doi.org/10.1080/14747731.2020.1807856>.
- [11] W. NORDHAUS, *Projections and uncertainties about climate change in an era of minimal climate policies,* Amer. Econ. J. Econ. Policy, 10 (2018), pp. 333–60, <https://doi.org/10.1257/pol.20170046>.
- [12] W. D. NORDHAUS, *Revisiting the social cost of carbon,* Proc. Natl. Acad. Sci. USA, 114 (2017), pp. 1518–1523, <https://doi.org/10.1073/pnas.1609244114>.
- [13] H. SCHIELZETH, *Simple means to improve the interpretability of regression coefficients,* Methods Ecol. Evol., 1 (2010), pp. 103–113, <https://doi.org/10.1111/j.2041-210X.2010.00012.x>.
- [14] I. M. SOBOL', D. ASOTSKY, A. KREININ, AND S. KUCHERENKO, *Construction and comparison of high-dimensional Sobol' generators,* Wilmott, 2011 (2011), pp. 64–79, <https://doi.org/10.1002/wilm.10056>.
- [15] K. SOETAERT, T. PETZOLDT, AND R. W. SETZER, *Solving differential equations in R: Package deSolve,* J. Statist. Softw., 33 (2010), pp. 1–25, <https://doi.org/10.18637/jss.v033.i09>.